

THE CAUSAL EFFECT OF FERTILITY ON COGNITIVE FUNCTIONING IN LATER LIFE

Journée de la Chaire Santé
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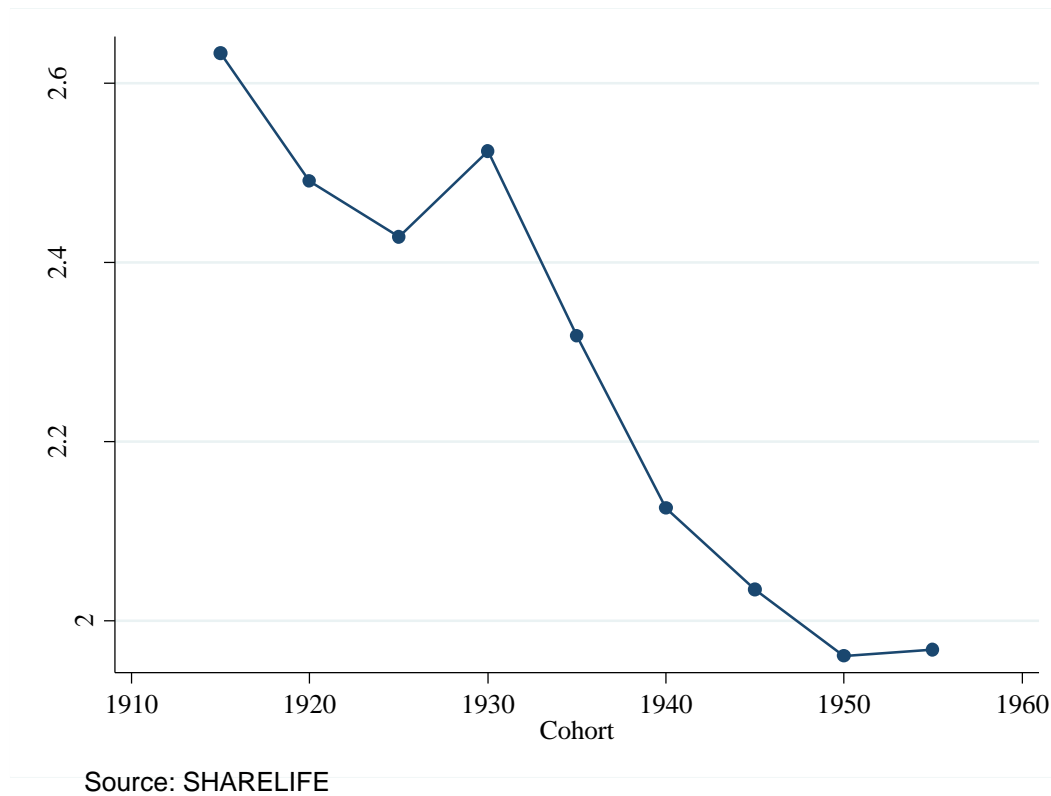
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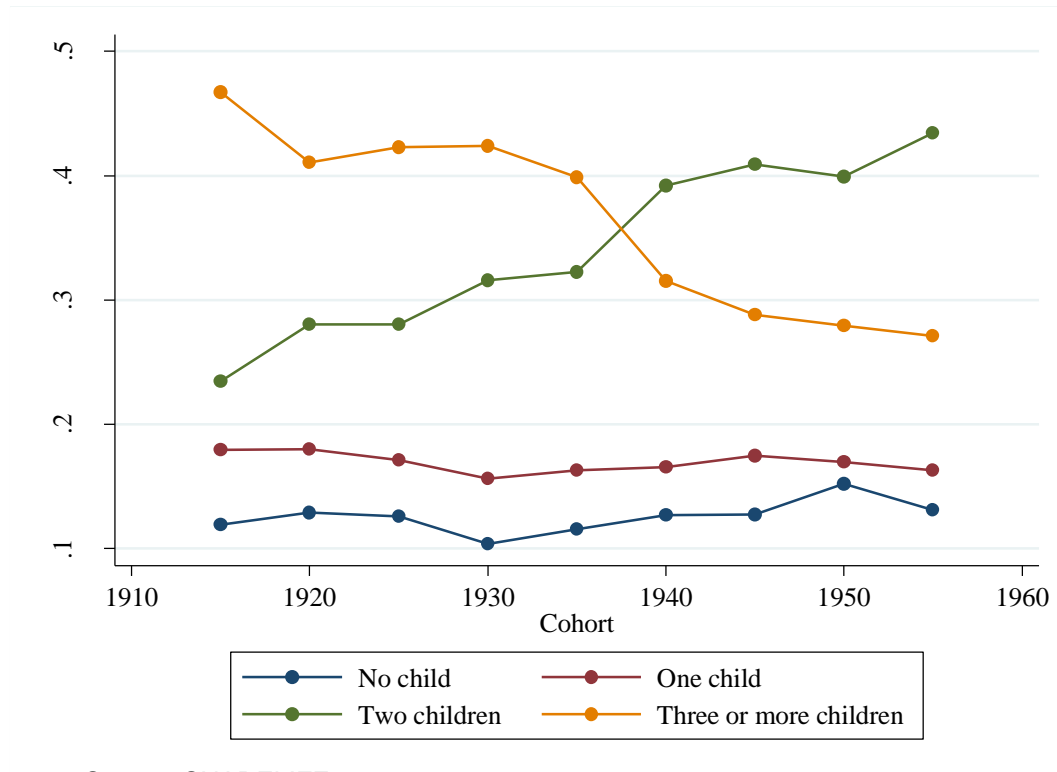
INTRODUCTION

- Ongoing demographic change implies an ageing of the population
- Childbearing patterns have changed dramatically over the last few decades
- What are the consequences of fertility changes on the health of older individuals?

INTRODUCTION: EVOLUTION OF PARITY ACROSS COHORT



INTRODUCTION: EVOLUTION OF PARITY ACROSS COHORT



Source: SHARELIFE

INTRODUCTION

- Older individuals face a variety of health challenges. Age-related decline in cognitive functioning is well documented (Baltes, Lindenberger, & Staudinger, 2006; Small, 2001)
- Cognitive impairments are associated with:
 - Loss of productivity
 - Decreased quality of life
 - Increased disability
 - Higher health-related expenditures

INTRODUCTION

- However this decline is not homogenous across the population, with some people maintaining cognitive vitality even into extreme old age (Berkman et al., 1993, Silver et al., 1998)
- This cognitive heterogeneity in ageing has been attributed by Stern and others (Stern, 2002, 2003; Scarmeas & Stern, 2003) to the concept of “cognitive reserve”
- The degree of resilience to biological changes has been found to depend on several factors, including socioeconomic dimensions such as education (Le Carret et al., 2003), occupation, leisure activities, or life style (see for a review: Fillit et al., 2002; and Fratiglioni et al., 2004)

INTRODUCTION

- The number of children may affect cognitive performance at older ages through several mechanisms:
 - Social:
 - Social contacts/support (Ertel et al 2008; Zunzunegui et al 2003; Vlachantoni et al 2015)
 - Economic:
 - Living standard (Mani et al. 2013; Caceres-Delpiano and Simonsen 2012)
 - Labour market participation (Rohwedder and Willis 2010; Bonsang et al 2012; Angrist and Evans 1998)
 - Health :
 - Lower estrogen exposure (Heys et al. 2011)
 - Higher risk of obesity (Caceres-Delpiano and Simonsen 2012)
 - Sleep deprivation (Ritcher et al. 2019; Virta et al. 2013)

THIS STUDY

- We contribute to the existing descriptive literature (e.g. Grundy and Read, 2017; Keenman and Grundy, 2018) by investigating the causal effect of fertility on cognitive functioning in later life
- We use an instrumental variable approach by exploiting a source of exogenous variation in the number of children
- The sex composition of the first two children is used as an instrument to study the marginal effect of additional children, conditional on having at least two children. There is a strong parental preference for a mixed sibling-sex composition (e.g. Ben-Porath and Welch, 1976)
- The use of this approach as an instrument was first proposed by Angrist and Evans (1998)

EMPIRICAL STRATEGY

- The equation to be estimated is the following:

$$cognition_{ict} = \beta_0 + \beta_1 children_{ict} + X'_{ict}\beta_2 + \alpha_c + \tau_t + \varepsilon_{ict}$$

- This equation can be estimated by OLS under the assumption that:

$$E(\varepsilon | children, X) = 0$$

- Unlikely to be the case due to selection, confounding factors or measurement error

EMPIRICAL STRATEGY

- Instrumental variables approach:

$$children_{ict} = \gamma_0 + \gamma_1 same_sex_{ict} + X'_{ict}\gamma_2 + \gamma_c + \delta_t + \omega_{ict}$$

$$cognition_{ict} = \beta_0 + \beta_1 \widehat{children}_{ict} + X'_{ict}\beta_2 + \alpha_c + \tau_t + \varepsilon_{ict}$$

- Hypotheses:

Instrument relevance

Instrument as good as random

Exclusion restriction

Monotonicity (in case of heterogeneous effects)

- Specification checks:

- Control for the sex of the first two children to discard any confounding bias to the potential effect of sex of children on cognitive functioning (Angrist and Evans, 1998)
- Alternative instrument using the fact of having two boys and the fact of having two girls. Allow to perform an overidentification test in order to see whether the sex of the children might bias the results (Angrist and Evans, 1998)

DATA: THE SAMPLE

- Pooled data from waves 1, 2, 4, 5 and 6 of the Survey of Health, Ageing and Retirement in Europe (SHARE)
- 19 European countries and Israel: Austria, Belgium, Croatia, Czech Republic, Denmark, Estonia, France, Germany, Greece, Ireland, Israel, Italy, Luxembourg, the Netherlands, Poland, Portugal, Slovenia, Spain, Sweden, and Switzerland
- Sample selection:
 - Individuals aged 65 year-old or more
 - Individuals with only biological children
 - Individuals with at least two children
- The final analytical sample includes 75,233 observations

DATA: THE MEASURES OF COGNITIVE FUNCTIONING

- The survey assesses cognitive functioning by using short and simple tests of episodic memory and executive functioning
- In the episodic memory task, participants were asked to memorize ten common words, and to list as many of these words as they could remember in one minute. There are an immediate and a delayed word recall task
- For the fluency task, respondents had to name as many different animals as possible in one minute
- We combine the score from each test using principal component analysis
- Results hold for each test separately

DATA: THE MEASURE OF PARITY

- The main independent variable is the number of biological children alive
- The instrument is based on a dummy variable equal to one when the two oldest children have the same sex

Sample distribution according to the sex of the first two children

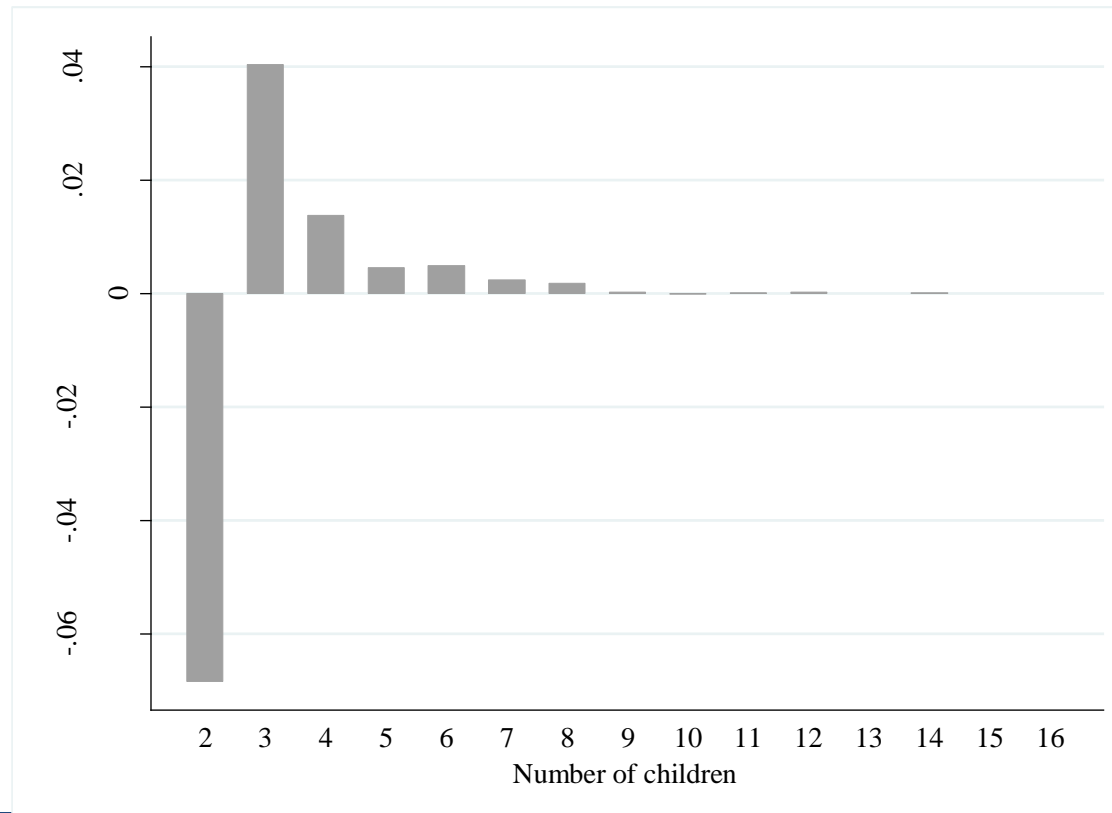
	Boy at second birth	Girl at second birth	Total
Boy at first birth	25.62%	25.13%	50.75%
Girl at first birth	24.95%	24.30%	49.25%
Total	50.57%	49.43%	100%

Probability to have at least three children according to the sex of the first two children

	Boy at second birth	Girl at second birth
Boy at first birth	48.0%	41.8%
Girl at first birth	41.7%	49.3%

DATA: THE MEASURE OF PARITY

- Difference in the distribution of the sample according to the sex composition of the first two children.



DATA: OTHER COVARIATES

- We control for a number of exogenous variables:
 - Country fixed effects
 - Sex country-specific effects
 - Wave fixed effects
 - Education
 - Age of the parent at second birth
 - Third-order polynomial in age
 - Being born abroad
- Results similar with or without controls (but gain in efficiency)

DESCRIPTIVE EVIDENCE

	All	By sex composition of the first two children		
		Different sex	Same sex	P-value for difference
Sex of the respondent (1=woman)	0.55	0.56	0.55	0.229
Age of the respondent	74.29	74.32	74.27	0.318
Age at the birth of the second child	29.44	29.46	29.42	0.205
Respondent born abroad	0.10	0.10	0.10	0.098
<u>Level of education :</u>				
Primary or less	0.35	0.35	0.35	0.989
Secondary	0.44	0.44	0.44	0.310
Tertiary	0.20	0.20	0.20	0.196
Education missing	0.01	0.01	0.01	0.814
Number of observations	75,233	37,676	37,557	

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Number of children	2.78	2.71	2.84	0.000***
Immediate word recall test	4.65	4.67	4.63	0.002***
Delayed word recall test	3.16	3.18	3.15	0.065*
Word fluency test	17.92	17.99	17.84	0.004***
Number of observations	75,233	37,676	37,557	

RESULTS: INSTRUMENT RELEVANCE, FIRST STAGE ESTIMATES: OLS

Dependent variable :	Number of children		
	All	Men	Women
Same sex	0.129*** (0.012)	0.128*** (0.018)	0.128*** (0.017)
Control variables included	yes	yes	yes
F-test of the excluded instrument	112.17	52.02	58.98
Two boys	0.113*** (0.015)	0.119*** (0.021)	0.106*** (0.021)
Two girls	0.145*** (0.016)	0.138*** (0.023)	0.151*** (0.021)
Control variables included	yes	yes	yes
F-test of the excluded instrument	57.19	26.02	31.24
P-value for difference between “Two boys” and “Two girls”	0.075	0.466	0.065
N	75,233	33,485	41,748

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: REDUCED FORM ESTIMATES, OLS

Dependent variable :	Cognitive test score		
	All	Men	Women
Same sex	-0.025*** (0.008)	-0.029** (0.011)	-0.021** (0.011)
Control variables included	yes	yes	yes
Two boys	-0.029*** (0.010)	-0.031** (0.014)	-0.028** (0.013)
Two girls	-0.021** (0.010)	-0.027* (0.014)	-0.015 (0.013)
Control variables included	yes	yes	yes
P-value for difference between “Two boys” and “Two girls”	0.459	0.814	0.419
N	75,233	33,485	41,748

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: 2SLS ESTIMATES

Dependent variable :	Cognitive test score index		
	All	Men	Women
Instrument used :	Same sex		
Number of children	-0.190*** (0.062)	-0.224** (0.093)	-0.164** (0.084)
Control variables included	yes	yes	yes
Endogeneity test (p-value)	0.006	0.023	0.101
Instruments used :	Two boys, two girls		
Number of children	-0.177*** (0.060)	-0.219** (0.092)	-0.139* (0.080)
Control variables included	yes	yes	yes
Endogeneity test (p-value)	0.011	0.025	0.176
Overidentification test (p-value)	0.200	0.611	0.207
N	75,233	33,485	41,748

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: INTERPRETATION

- The effect of having one additional child is equivalent to a cognitive ageing of about 4 years of age.
- The IV results are larger than OLS estimates. It suggests positive selection into those who decide to have more than two children.*
- Does not estimate an average effect of the population. The 2SLS results estimate a Local Average Treatment Effect (LATE).
- Compliers are more likely to be:
 - Less educated
 - Had their second child at younger age
- But no difference across:
 - countries/regions
 - age
 - sex
 - immigrant status

RESULTS: 2SLS ESTIMATES BY LEVEL OF EDUCATION OF THE PARENTS

Dependent variable :	Cognitive test score index		
	Primary education	Secondary education	Tertiary education
Instrument used :	Same sex		
Number of children	-0.187** (0.077)	-0.198* (0.113)	-0.131 (0.170)
Control variables included	yes	yes	yes
F-test of excluded instruments	45.3	47.1	21.4
N	26,476	33,197	15,066

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: 2SLS ESTIMATES BY AGE AT SECOND BIRTH

Dependent variable :	Cognitive test score index		
	All	Men	Women
Number of children	-0.212*** (0.072)	-0.240** (0.095)	-0.271** (0.126)
Number of children x Age at second birth	-0.017 (0.017)	0.022 (0.036)	-0.039* (0.022)
Control variables included	yes	yes	yes
N	75,233	33,485	41,748

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: HETEROGENEITY

	Cognitive test score index			
	Northern Europe	Western Europe	Eastern Europe	Southern Europe
Number of children	-0.535** (0.244)	-0.150 (0.123)	-0.142 (0.120)	-0.097 (0.095)
Control variables included	yes	yes	yes	yes
Endogeneity test (p-value)	0.006	0.308	0.402	0.638
	First stage estimates: Number of children			
Same sex	0.118*** (0.029)	0.109*** (0.021)	0.134*** (0.018)	0.149*** (0.028)
Control variables included	yes	yes	yes	yes
F-test of the excluded instruments	16.269	26.943	54.829	27.807
N	9,326	29,443	16,945	17,189

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: HETEROGENEITY

	Cognitive test score index				
	Northern Europe	Western Europe	Eastern Europe	Southern Europe	All but not Northern countries
Number of children	-0.535** (0.244)	-0.150 (0.123)	-0.142 (0.120)	-0.097 (0.095)	-0.132** (0.067)
Control variables included	yes	yes	yes	yes	yes
Endogeneity test (p-value)	0.006	0.308	0.402	0.638	0.147
	First stage estimates: Number of children				
Same sex	0.118*** (0.029)	0.109*** (0.021)	0.134*** (0.018)	0.149*** (0.028)	0.127*** (0.013)
Control variables included	yes	yes	yes	yes	yes
F-test of the excluded instruments	16.269	26.943	54.829	27.807	91.465
N	9,326	29,443	16,945	17,189	63,577

Note. Clustered (at the individual level) standard errors are in parentheses. * $p < .1$. ** $p < .05$. *** $p < .01$.

RESULTS: MECHANISMS

- **Social support/contacts:**
 - Positive with the frequency of contacts (smaller effect in the North, larger in the South)[*](#)
 - Not significant with informal care (no difference across regions)[*](#)
 - Positive with the number of grandchildren (larger in the North, smaller effect in the South)[*](#)
 - Not significant with caring for grandchildren (no difference across regions)[*](#)
- **Labour market experience:**
 - Not significant regarding the age they quitted their last job (no difference across regions)[*](#)
 - Not significant regarding the type of profession (difference across regions)[*](#)
- **Standard of living:**
 - Negative for wealth (Larger in Northern and Western Europe)[*](#)
 - Not significant for income (Negative for Northern Europe)[*](#)
- **Health:**
 - Positive for BMI and the risk of being overweight[*](#)

INTERPRETATIONS AND LIMITATIONS

- Given the nature of the instruments, it only identifies the effect of having a third child.
- LATE
- Loss of precision due to the use of 2SLS. Difficulties in exploring the heterogeneity of the effects.

CONCLUSIONS

- Using an instrumental variable approach, we estimated the causal effect of having a third child on cognitive functioning in later life in Europe
- We find that it has large negative effect on the level of cognitive performance in later life for both men and women
- This result suggests that the trend toward having less children could have unexpected benefits for cognition in later life
- More research is warranted in order to better understand the mechanisms linking childbearing history and cognitive functioning in later life